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STATISTICAL INFERENCE FOR PARAMETERS OF GOMPERTZ DISTRIBUTION BASED ON GENERAL PROGRESSIVELY TYPE-II RIGHT CENSORED ORDER STATISTICS

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ABSTRACT: In this article, we concluded the Bayesian estimations for the parameters of the Gompertz distribution (GD) based on general progressively type-II right censored order statistics (GPTIIRCO). The squared error loss function (SEL) and linear-exponential loss function (LINEX) and the generalization of the entropy loss function (GE) and Al-Bayyati loss function (ABL) are considered for Bayesian estimation. Finally, a numerical example is established to clear the theoretical procedures.

KEYWORDS: Gompertz Distribution; General Progressively type-II Right Censored Order Statistics.

INTRODUCTION

Kim [11] derived on estimating Burr Type XII parameter based on GPTIICO. Mohie El-Din et al. [12] derived Estimation for parameters of Feller-Pareto distribution from progressive Type-II censoring and some characterizations. Mohie El-Din et al. [13] and [14] derived characterization for Gompertz and linear failure rate distributions using recurrence relations of single and product moments based on GPTIICO.

In failure data analysis, it is common that some individuals cannot be observed for the full failure times. GPTIICOS is a useful and more general scheme in which a specific fraction of individuals at risk may be removed from the study at each of several ordered failure times. Progressively censored samples have been considered, among others, as solved by Davis and Feldstein [7], Balakrishnan et al. [4], and Guilbaud [10]. This scheme of censoring was generalized by Balakrishnan and Sandhu [3] as follows: at time $X_0 \equiv 0$, nunits are placed on test; the first r failure times, $X_1, ..., X_r$, are not observed; at time $X_i + 0$, where X_i is the *i*th ordered failure time (i = r + 1, ..., m - 1), R_i units are removed from the test randomly, so prior to the $(i + 1)^{th}$ failure there are $n_i = n - i - \sum_{j=r+1}^{i} R_j$ units on test; finally, at the time of the m^{th} failure, X_m , the experiment is terminated, i.e., the remaining R_m units are removed from the test. The R_i 's, m and r are prespecified integers which must satisfy the conditions: $0 \le r < m \le n$, $0 \le R_i \le n_{i-1}$ for i = r + 1, ..., m - 1 with $n_r = n - r$ and $R_m = n_{m-1} - 1$.

If the failure times are based on an absolutely continuous distribution function (cdf) F with probability density function (pdf) f, the joint probability density function based on GPTIICOS failure times $X_{r+1:m:n}, X_{r+2:m:n}, \dots, X_{m:m:n}$, is given by

$$f_{X_{r+1:m:,n},X_{r+2:m:,n},\dots,X_{m:m:,n}}(x_{r+1},x_{r+2},\dots,x_m) = A_{(n,m-1)}[F(x_{r+1},\theta)]^r \times \prod_{i=r+1}^m f(x_i,\theta)[1-F(x_i,\theta)]^{R_i}, \qquad x_{r+1} < x_{r+2} < \dots < x_m,$$
(1.1)

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$$A_{(n,m-1)} = \frac{n!}{r! (n-r)!} \left(\prod_{j=r}^{m-1} n_j \right), \quad n_i = n - i - \sum_{j=r+1}^i R_j , \quad i = r+1, \dots, m-1.$$

Through ut this paper, we concluded the Bayesian estimations for the parameters of the GD based on GPTIIRCO. The SEL, ABL, LINEX and GE loss functions are considered for Bayesian estimation. Finally, a numerical example is established to clear the theoretical procedures are obtained. For more details of the SEL, ABL, LINEX and GE loss functions see Al-Aboud [1], Al-Bayyati [2], Bernardo and Smith [5], Calabria and Pulcini [6], and Figueiredo [8].

Let $X_{r+1:m:n}^{(R_{r+1},R_{r+2},...,R_m)} < X_{r+2:m:n}^{(R_{r+1},R_{r+2},...,R_m)} < \cdots < X_{m:m:n}^{(R_{r+1},R_{r+2},...,R_m)}$ be the *m* ordered observed failure times in a sample of size (n - r) under GPTIICOS scheme from the GD with probability density function (pdf) given by

$$f(x,\alpha,\beta) = \alpha e^{\beta x - \frac{\alpha}{\beta}(e^{\beta x} - 1)}, \quad \alpha > 0, \quad \beta > 0 \qquad x \ge 0$$

$$(1.2)$$

and the corresponding cumulative distribution function (cdf) is given by

$$F(x, \alpha, \beta) = 1 - e^{-\frac{\alpha}{\beta}(e^{\beta x} - 1)}, \quad \alpha > 0, \ \beta > 0 \quad x \ge 0.$$
(1.3)

The hazard function at time *t* is given, by

$$H(t) = \frac{f(t)}{F(t)} = \alpha e^{\beta t}.$$
(1.4)

The GD was introduced by Gompertz [9].

Bayes estimation

In this section, we have three types of loss functions, which are SEL, ABL, LINEX and GE loss functions. We apply these types in two cases, the first is known shape parameter β and the second is

unknown shape β and scale parameter.

The likelihood function is

$$L(\alpha,\beta|\mathbf{x}) = A_{(n,m-1)} \left[1 - e^{-\frac{\alpha}{\beta} \left(e^{\beta x_{r+1}} - 1 \right)} \right]^r \prod_{i=r+1}^m \alpha e^{\beta x_i - \frac{\alpha}{\beta} \left(e^{\beta x_i} - 1 \right)} \left[e^{-\frac{\alpha}{\beta} \left(e^{\beta x_i} - 1 \right)} \right]^{R_i}.$$
 (2.1)

The case of known shape parameter β

For Bayesian estimation, we use here the natural conjugate prior density function for α and the known shape parameter β is given by

$$\pi(\alpha | \mathbf{x}, \beta) = \frac{b^a}{\Gamma(a)} \alpha^{a-1} e^{-b\alpha}, \quad a, b > 0,$$
(2.2)
Combining the likelihood function given by Eq. (2.1) and prior function Eq. (2.2), the posterior density of π

likelihood function given by Eq.(2.1) and prior function Eq.(2.2), the posterior density of α given x is

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$$\pi^{*}(\alpha | \mathbf{x}, \beta) = \frac{L(\alpha | \mathbf{x}, \beta) \pi(\alpha | \mathbf{x}, \beta)}{\int_{0}^{\infty} L(\alpha | \mathbf{x}, \beta) \pi(\alpha | \mathbf{x}, \beta) d\alpha}$$
$$= \frac{L(\alpha | \mathbf{x}, \beta) \pi(\alpha | \mathbf{x}, \beta)}{I(a, b)} .$$
(2.3)

Bayes estimator based on squared error loss function (SEL) The Bayes estimators $\hat{\alpha}_{BS}$ of α based on SEL loss function given by

$$\hat{a}_{BS} = E(\alpha | \mathbf{x}, \beta)$$

$$= \int_{0}^{\infty} \alpha \pi^{*}(\alpha | \mathbf{x}, \beta) d\alpha$$

$$= \frac{a \, l(a+1, b)}{b \, l(a, b)}.$$
(2.4)

The hazard function H(t) at time under the SEL loss function given by

$$H_{BS}(t) = E(\Pi(t)|\mathbf{x},\beta)$$
$$= e^{t\beta}\hat{\alpha}_{BS}.$$
 (2.5)

Bayes estimator based on Al-Bayyati loss function (ABL) The Bayes estimators $\hat{\alpha}_{BA}$ of α based on ABL loss function given by

$$\hat{\alpha}_{BA} = \frac{E(\alpha^{q+1}|\mathbf{x},\beta)}{E(\alpha^{q}|\mathbf{x},\beta)}$$

$$= \frac{\int_{0}^{\infty} \alpha^{q+1} \pi^{*}(\alpha|\mathbf{x},\beta) d\alpha}{\int_{0}^{\infty} \alpha^{q} \pi^{*}(\alpha|\mathbf{x},\beta) d\alpha}$$

$$= \frac{(a+q) I(a+q+1,b)}{b I(a+q,b)}.$$
(2.6)

Similarly, hazard function H(t) at time t under the ABL loss function given by

$$\hat{H}_{BA}(t) = \frac{E(\mathrm{H}^{q+1}(t)|\mathbf{x},\beta)}{E(\mathrm{H}^{q}(t)|\mathbf{x},\beta)}$$
$$= e^{t\beta} \hat{\alpha}_{BA}.$$
(2.7)

Bayes estimator based on linear-exponential loss function (LINEX)

The Bayes estimators $\hat{\alpha}_{BL}$ of α based on **LINEX** loss function given by

$$\hat{\alpha}_{BL} = \frac{-1}{p} \log \left(E(e^{-p\alpha} | \mathbf{x}, \beta) \right)$$
$$= \frac{-1}{p} \log \left(\int_0^\infty e^{-p\alpha} \pi^*(\alpha | \mathbf{x}, \beta) d\alpha \right)$$

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$$= \frac{-1}{p} \log \left(\frac{b^{a} \ I(a, b+p)}{(b+p)^{a} \ I(a, b)} \right).$$
(2.8)

Similarly, hazard function H(t) at time t under the LINEX loss function given by

$$\begin{split} \hat{H}_{BA}(t) &= \frac{-1}{p} \log \left(E\left(e^{-pH(t)} | \boldsymbol{x}, \beta\right) \right) \\ &= \frac{-1}{p} \log \left(\frac{\mathbf{b}^a \ \mathbf{I}\left(\mathbf{a}, \mathbf{b} + \mathbf{p} \ e^{t\beta}\right)}{(\mathbf{b} + \mathbf{p} \ e^{t\beta})^a \ \mathbf{I}\left(\mathbf{a}, \mathbf{b}\right)} \right). \end{split}$$
(2.9)

Bayes estimator based ongeneralization of the entropy loss function (GE)

The Bayes estimators $\hat{\alpha}_{BG}$ of α based on GE loss function given by

$$\hat{\alpha}_{BG} = \left(E(\alpha^{-v} | \mathbf{x}, \beta) \right)^{\frac{-v}{v}} = \left(\int_{0}^{\infty} \alpha^{-v} \pi^{*}(\alpha | \mathbf{x}, \beta) d\alpha \right)^{\frac{-1}{v}} = \left(\frac{b^{v} \Gamma(a - v) I(a - v, b)}{\Gamma(a) I(a, b)} \right)^{\frac{-1}{v}}.$$
(2.10)

The hazard function H(t) at time t under the GE loss function given by

$$\hat{H}_{BG}(t) = \left(E(\mathbf{H}^{-\nu}(t)|\mathbf{x},\beta) \right)^{\frac{-1}{\nu}}$$
$$= e^{t\beta} \,\hat{\alpha}_{BS} \,. \tag{2.11}$$

The case of unknown shape and scale parameters α and β

For Bayesian estimation, we use here the natural conjugate prior density function for (α, β) given by

$$\pi(\alpha, \beta | \mathbf{x}) = \pi_1(\beta) \pi_2(\alpha | \beta)$$

$$=\frac{b_1^{a_1}b_2^{a_2}}{\Gamma(a_1)\Gamma(a_2)}\alpha^{a_2-1}\beta^{a_1-a_2-1}e^{-\frac{b_2\alpha}{\beta}-b_1\beta}, \quad a_1,a_2,b_1,b_2>0,$$
(2.12)

where

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 $\pi_{1}(\beta) = \frac{b_{1}^{a_{1}}}{\Gamma(a_{1})}\beta^{a_{1}-1}e^{-b_{1}\beta} and \pi_{2}(\alpha|\beta) = \frac{(\overline{\beta})^{-1}}{\Gamma(a_{2})}\alpha^{a_{2}-1}e^{-\frac{b_{2}\alpha}{\beta}}.$

Combining the likelihood function given by Eq. (2.1) and prior function Eq. (2.12), the posterior

density of α and β given x is

$$\pi^{*}(\alpha,\beta|\mathbf{x}) = \frac{L(\alpha,\beta|\mathbf{x})\pi(\alpha,\beta|\mathbf{x})}{\int_{0}^{\infty}\int_{0}^{\infty}L(\alpha,\beta|\mathbf{x})\pi(\alpha,\beta|\mathbf{x})d\alpha \ d\beta}$$

Published by European Centre for Research Training and Development UK (www.eajournals.org) $= \frac{L(\alpha, \beta | \mathbf{x}) \pi(\alpha, \beta | \mathbf{x})}{I(a_1, b_1, a_2, b_2)} . \qquad (2.13)$

Bayes estimator based on squared error loss function (SEL)

The Bayes estimators α and β based on SEL loss function are given respectively, by

$$\hat{\alpha}_{BS} = E(\alpha | \mathbf{x}) = \int_{0}^{\infty} \int_{0}^{\infty} \alpha \pi^{*}(\alpha, \beta | \mathbf{x}) d\alpha \ d\beta$$
$$= \frac{a_{1} a_{2} I(a_{1} + 1, b_{1}, a_{2} + 1, b_{2})}{b_{1} b_{2} I(a_{1}, b_{1}, a_{2}, b_{2})},$$
(2.14)

and

$$\hat{\beta}_{BS} = E(\beta | \mathbf{x})$$

$$= \frac{a_1 \ \mathbf{I}(a_1 + 1, b_1, a_2, b_2)}{b_1 \ \mathbf{I}(a_1, b_1, a_2, b_2)}.$$
(2.15)

The hazard function H(t) at time under the SEL loss function given by

$$\hat{H}_{BS}(\mathbf{t}) = E(II(\mathbf{t})|\mathbf{x}) = \frac{a_1 a_2 b_1^{a_1}}{b_1 b_2 (b_1 - t)^{a_1}} \frac{I(a_1 + 1, b_1 - t, a_2 + 1, b_2)}{I(a_1, b_1, a_2, b_2)}.$$
(2.16)

Bayes estimator based on Al-Bayyati loss function (ABL) The Bayes estimators α and β based on ABL loss function given by

$$\hat{\alpha}_{BA} = \frac{E(\alpha^{q+1}|\mathbf{x})}{E(\alpha^{q}|\mathbf{x})}$$
$$= \frac{\int_{0}^{\infty} \int_{0}^{\infty} \alpha^{q+1} \pi^{*}(\alpha|\mathbf{x}) d\alpha \, d\beta}{\int_{0}^{\infty} \int_{0}^{\infty} \alpha^{q} \pi^{*}(\alpha|\mathbf{x}) d\alpha \, d\beta}$$
$$= \frac{(a+q) \, l(a+q+1,b)}{b \, l(a+q,b)}, \tag{2.17}$$

and

$$\hat{\beta}_{BS} = \frac{a_1 \ \mathsf{I}(a_1 + 1, b_1, a_2, b_2)}{b_1 \ \mathsf{I}(a_1, b_1, a_2, b_2)}.$$
(2.18)

Similarly, hazard function H(t) at time t under the ABL loss function given by

$$\hat{H}_{BA}(t) = \frac{E(\mathrm{H}^{q+1}(t)|\mathbf{x})}{E(\mathrm{H}^{q}(t)|\mathbf{x})}$$
$$= e^{t\beta} \hat{\alpha}_{BA}.$$
(2.19)

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Bayes estimator based on linear-exponential loss function (LINEX)

The Bayes estimators α and β based on **LINEX** loss function given by

$$\hat{\alpha}_{BL} = \frac{-1}{p} \log \left(E(e^{-p\alpha} | \mathbf{x}) \right)$$
$$= \frac{-1}{p} \log \left(\int_{0}^{\infty} \int_{0}^{\infty} e^{-p\alpha} \pi^{*}(\alpha | \mathbf{x}) d\alpha \ d\beta \right)$$
$$= \frac{-1}{p} \log \left(\frac{b_{2}^{a_{2}} \ I(a_{1}, b_{1}, a_{2}, b_{2} + p)}{(b_{2} + p)^{a_{2}} \ I(a_{1}, b_{1}, a_{2}, b_{2})} \right),$$
(2.20)

and

$$\hat{\beta}_{BL} = \frac{-1}{p} \log \left(E(e^{-p\beta} | \mathbf{x}) \right)$$
$$= \frac{-1}{p} \log \left(\frac{b_1^{a_1} \ I(a_1, b_1 + p, a_2, b_2)}{(b_1 + p)^{a_1} \ I(a_1, b_1, a_2, b_2)} \right).$$
(2.21)

Similarly, hazard function H(t) at time t under the LINEX loss function given by

$$\hat{H}_{BA}(t) = \frac{-1}{p} \log \left(E\left(e^{-pH(t)} | \mathbf{x}\right) \right)$$
$$= \frac{-1}{p} \log \left(\sum_{l=0}^{\infty} \frac{(-p)^l b_1^{a_1} \Gamma(a_1 + l) \Gamma(a_2 + l) I(a_1 + l, b_1 - tl, a_2 + l, b_2)}{\Gamma(a_1) \Gamma(a_2) \Gamma(l + 1) (b - tl)^{a_1} I(a_1, b_1, a_2, b_2)} \right).$$
(2.22)

Bayes estimator based ongeneralization of the entropy loss function (GE)

The Bayes estimators α and β based on GE loss function given by

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$$\hat{\alpha}_{BG} = \left(E(\alpha^{-\nu} | \mathbf{x}) \right)^{\overline{\nu}}$$

$$= \left(\int_{0}^{\infty} \int_{0}^{\infty} \alpha^{-\nu} \pi^{*}(\alpha | \mathbf{x}) d\alpha \, d\beta\right)^{\frac{-1}{\nu}}$$
$$= \left(\frac{(b_{1} \ b_{2})^{\nu} \Gamma(a_{1} - \nu) \Gamma(a_{2} - \nu) I(a_{1} - \nu, b_{1}, a_{2} - \nu, b_{2})}{\Gamma(a_{1}) \Gamma(a_{2}) I(a_{1}, b_{1}, a_{2}, b_{2})}\right)^{\frac{-1}{\nu}},$$
(2.23)

and

-1

$$\hat{\beta}_{BG} = \left(E(\beta^{-\nu} | \mathbf{x}) \right)^{\frac{\nu}{\nu}} = \left(\frac{(b_1)^{\nu} \Gamma(a_1 - \nu) I(a_1 - \nu, b_1, a_2, b_2)}{\Gamma(a_1) I(a_1, b_1, a_2, b_2)} \right)^{\frac{-1}{\nu}}.$$
(2.24)

The hazard function H(t) at time t under the GE loss function given by

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$$\hat{H}_{BG}(t) = \left(E(\mathbf{H}^{-\nu}(t)|\mathbf{x}) \right)^{\frac{1}{\nu}} = \left(\frac{(b_1)^{a_1+\nu} (b_2)^{\nu} \Gamma(a_1-\nu) \Gamma(a_2-\nu) I(a_1-\nu, b_1+t\nu, a_2-\nu, b_2)}{\Gamma(a_1) \Gamma(a_1) (b_1+t\nu)^{a_1} I(a_1, b_1, a_2, b_2)} \right)^{\frac{-1}{\nu}}.$$
(2.25)

Numerical results

In this section, we obtained some numerical results to compare the Bayes estimation under SEL, ABL

(with q = 0.5 and q = -0.5), LINEX (with p = 0.5) and GE (with v = -0.5) loss functions according to the following steps:

Algorithm

- Specify the values of n, m, r, t, q, p and v.
- 2- Specify the values of the parameters α and β .
- 3- For given values of the prior parameters.
- 4- Generate a simple random sample of size *m* from U(0,1) distribution, $(U_{r+1}, ..., U_m)$.
- 5- Determine the values of the censoring scheme R_{i} , i = r + 1, ..., m.

6- Set
$$E_i = U_i^{\frac{1}{\sum_{j=m-i+1}^m R_j}}, \quad i = r+1, ..., m.$$

7- Obtain the General Progressively type-II censored sample $(U_{r+1:m:n}^*, ..., U_{m:m:n}^*)$, where $U_{i:m:n}^* = 1 - \prod_{j=m-i+1}^m E_i$, i = r+1, ..., m.

8- From steps (4)-(7), the order observations $y_{r+1:m:n}, y_{r+2:m:n}, \dots, y_{n_1:m:n}, y_{n_1+1:m:n}, \dots, y_{m:m:n}$ are calculated as follows

$$y_{i:m:n} = \frac{1}{\beta} \log \left[1 - \frac{\beta}{\alpha} \log(1 - U_{i:m:n}^*) \right], \ i = r + 1, ..., m.$$

9- We perform Monte Carlo simulation to compare th performaces of the different estimators for different sampling schemes. Monte Carlo simulations were performed utilizing 1000 general progressively type-II censored samples for each simulations. The mean squared error (MSE) is used to compear the estimators, where $MSE = \frac{1}{1000} \sum_{i=1}^{1000} (\theta - \hat{\theta})^2$ and $\hat{\theta}$ is the estimator of θ .

10- Compute the BEs of the model parameters relative to SEL, ABL, LINEXL and GE loss functions.

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Sch	neme n		m	r	R
1	30	23	2		-,-, 0,0,0,0,0,1,0,0,0,0,1,0,1,0,0,0,1,1,1,1
2	30	23	5		-,-, -,-,0,0,1,0,0,0,0,0,1,0,1,0,0,0,0,1,1,1,1
3	30	17	2		-,-, 0,0,0,0, 1,1,1,1,1,1,1,2,2,1,1
4	30	17	5		-,-, -,-,0, 1,1,1,1,1,1,2,2,1,1
5	10	07	1		-,0,0,0,1,1,1

The following is the generated sample from Gompertz distribution pdf in (1).

The case of known parameter β

Table 1: The values of BIAS and MSE of the Bayesian estimates of α based on SEL, ABL, LINEX and GE loss functions when $\beta = 0.3$ and 0.5.

n	n m		α	β	\hat{a}_{BS}	<u> </u>	l = -0.5	$\hat{\alpha}_{BA}$	q=0.5	\hat{a}_{BI}	<i>.</i>	$\hat{\alpha}_{BE}$	·
п	m	r	u	Ρ	BIAS M.	SE BIAS	MSE	BIAS	MSE	BIAS	MSE	BIAS	MSE
30	23	2	0.3	0.3	0.01511 0.0	0517 0.00829	0.00480	0.02193	0.00564	0.01399	0.00506	0.01172	0.0049
				0.5	0.01947 0.0	0522 0.01256	0.00479	0.02638	0.00575	0.01832	0.00510	0.01603	0.0050
			0.5	0.3	0.02669 0.0	1381 0.01530	0.01277	0.03809	0.01512	0.02358	0.01333	0.02103	0.0132
		5	0.3	0.3	0.02091 0.0	0590 0.01396	0.00543	0.02786	0.00648	0.01974	0.00577	0.01745	0.0056
			0.5	0.5 0.3	0.01025 0.0 0.02358 0.0	0431 0.00354 1330 0.01224	$\begin{array}{c} 0.00404 \\ 0.01234 \end{array}$	0.01697 0.03492	0.00468 0.01452	0.00917 0.02050	0.00423 0.01285	0.00691 0.01794	0.0041(0.0127
	17	2	0.3	0.3	0.01873 0.0	0701 0.00941	0.00637	0.02804	0.00784	0.01715	0.00682	0.01410	0.0066
				0.5	0.01827 0.0	0703 0.00897	0.00639	0.02758	0.00785	0.01671	0.00684	0.01366	0.0066
			0.5	0.3	0.03043 0.0	1808 0.01493	0.01639	0.04593	0.02028	0.02612	0.01727	0.02274	0.0171
		5	0.3	0.3	0.02085 0.0	0671 0.01147	0.00604	0.03024	0.00756	0.01927	0.00652	0.01620	0.0063
				0.5	0.02093 0.0	0714 0.01154	0.00645	0.03032	0.00802	0.01934	0.00694	0.01627	0.0067
			0.5	0.3	0.03357 0.0	1981 0.01796	0.01793	0.04918	0.02221	0.02918	0.01889	0.02582	0.0188
10) 7	1	0.3	0.3	0.05179 0.0	2691 0.02703	0.02167	0.07654	0.03362	0.04670	0.02483	0.03963	0.0241
				0.5	0.05211 0.0	2489 0.02734	0.01991	0.07689	0.03132	0.04709	0.02300	0.03995	0.0222
			0.5	0.3	0.08072 0.0	6527 0.03986	0.05236	0.12159	0.08209	0.06729	0.05746	0.06066	0.0584

Table 2 : The values of BIAS and MSE of the Bayesian estimates of $H(t)$ when $t = 3$ based on
SEL, ABL, LINEX and GE loss functions when $\beta = 0.3$ and 0.5.

n	m	r	α	β	HBS BIAS	s MSE	Ĥ _{BA} q BIAS	=-0.5 MSE	H _{BA} BIAS	q=0.5 MSE	$\frac{H_{B}}{BIAS}$	L MSE	H _{BI}	E MSE
					DIND	MDL	DIND	MDL	DIND	MDL	DIND	1452	DIND	1152
30	23	2	0.3	0.3	0.03717	0.03127	0.02040	0.02902	0.05394	0.03411	0.03043	0.02970	0.02883	0.0300
				0.5	0.08725	0.10487	0.05627	0.09626	0.11823	0.11549	0.06453	0.09497	0.07184	0.1003
			0.5	0.3	0.06566	0.08354	0.03763	0.07726	0.09369	0.09148	0.04703	0.07680	0.05172	0.0802
		5	0.3	0.3	0.05143	0.03571	0.03433	0.03283	0.06852	0.03921	0.04441	0.03381	0.04293	0.0342
				0.5	0.04596	0.08660	0.01584	0.08112	0.07607	0.09398	0.02457	0.07982	0.03098	0.0836
	17	2	0.5 0.3	$^{0.3}_{0.3}$	$\begin{array}{c} 0.05800 \\ 0.04606 \end{array}$	$\begin{array}{c} 0.08043 \\ 0.04243 \end{array}$	0.03011 0.02315	0.07467 0.03852	0.08589 0.06897	0.08782 0.04746	$\begin{array}{c} 0.03958 \\ 0.03665 \end{array}$	$\begin{array}{c} 0.07419 \\ 0.03968 \end{array}$	$0.04413 \\ 0.03469$	0.0773
				0.5	0.08190	0.14121	0.04021	0.12837	0.12359	0.15775	0.05117	0.12543	0.06121	0.1343

Print ISSN: ISSN 2055-0154(Print), Online ISSN: ISSN 2055-0162(Online)

European Journal of Statistics and Probability

Vol.5, No.3, pp.41-51, June 2017

_Published by European Centre for Research Training and Development UK (www.eajournals.org)

 0.5
 0.3
 0.07485
 0.10936
 0.03672
 0.09913
 0.11298
 0.12267
 0.04919
 0.09805
 0.05593
 0.10384

 5
 0.3
 0.05130
 0.04057
 0.02821
 0.03655
 0.07439
 0.04573
 0.04179
 0.03784
 0.03984
 0.03843

 0.5
 0.09382
 0.14343
 0.05173
 0.12954
 0.13590
 0.16109
 0.06257
 0.12655
 0.07293
 0.13606

 10
 7
 1
 0.3
 0.12737
 0.16279
 0.13436
 0.05647
 0.10701
 0.06351
 0.11379

 10
 7
 1
 0.3
 0.12737
 0.16279
 0.18827
 0.20337
 0.09748
 0.14612

 0.5
 0.23355
 0.49993
 0.12251
 0.39992
 0.34461
 0.62902
 0.13973
 0.3592
 0.17904
 0.44720

 0.5
 0.3
 0.19855
 0.39484
 0.31677
 0.29907
 0.49664
 0.12127
 0.29300
 0.14920
 0.35355

</tabu/>

The case of unknown parameters α and β

Table 3: The values of BIAS and MSE of the Bayesian estimates of α based on SEL, ABL, LINEX and GE loss functions

 n	m	r	α	β	- a _E	15	$\hat{\alpha}_{BA}$	r = -0.5	α_{BA}	q=0.5	a _B	L	α _B	5
			u	Ρ	BIAS	MSE	BIAS	MSE	BIAS	MSE	BIAS	MSE	BIAS	MSE
30	23	2	0.3	0.3	0.11413	0.02057	0.10177	0.01748	0.12651	0.02398	0.11149	0.01977	0.10799	0.01899
				0.5	0.19652	0.04644	0.18184	0.04045	0.21121	0.05288	0.19280	0.04476	0.18923	0.04341
			0.5	0.3	0.12476	0.03210	0.10639	0.02691	0.14314	0.03799	0.11886	0.03002	0.11564	0.02944
		5	0.3	0.3	0.11381	0.02017	0.10145	0.01711	0.12617	0.02356	0.11117	0.01939	0.10767	0.01861
				0.5	0.19663	0.04612	0.18195	0.04014	0.21132	0.05254	0.19292	0.04444	0.18934	0.04309
			0.5	0.3	0.11667	0.03107	0.09853	0.02617	0.13482	0.03665	0.11090	0.02908	0.10767	0.02855
	17	2	0.3	0.3	0.08188	0.01364	0.07062	0.01153	0.09314	0.01602	0.07964	0.01311	0.07629	0.01256
				0.5	0.13706	0.02658	0.12422	0.02278	0.14990	0.03073	0.13417	0.02558	0.13069	0.02466
		5	0.5 0.3	$5 & 0.3 \\ $	0,08994 0.07655	0.02588 0.01163	0.07267 0.06544	0.02205 0.00972	0.1072 0.08766	1 0.03034 0.01379	0.08465 0.07439	0.02430 0.01117	0.08132 0.07104	7 0.0239 0.01065
				0.5	0.13512	0.02560	0.12234	0.02189	0.14791	0.02965	0.13226	0.02464	0.12878	0.02372
			0.5	0.3	0.08037	0.02223	0.06337	0.01888	0.09736	0.02618	0.07527	0.02087	0.07193	0.02049
10	7	1	0.3	0.3	0.10079	0.02819	0.07254	0.02084	0.12906	0.03731	0.09463	0.02576	0.08692	0.02436
				0.5	0.16657	0.05056	0.13370	0.03756	0.19945	0.06596	0.15832	0.04586	0.15043	0.04388
			0.5	0.3	0.10816	0.05879	0.06526	0.04483	0.15108	0.07694	0.09394	0.05083	0.08709	0.05142

Table 4 : The values of BIAS and MSE of the Bayesian estimates of β based on SEL, ABL,
LINEX and GE loss functions

 n	m	r	α	β	$\frac{\overline{\beta}_{BS}}{BIAS MSE}$	β _{BA} BIAS	q=-0.5 MSE	β _{BA} BIAS	q=0.5 MSE	$\frac{\overline{\beta}_B}{BIAS}$	L MSE	β _{BI} BIAS	E MSE
30	23	2	0.3	0.3	0.19074 0.03639	0.19566	0.03830	0.18584	0.03455	0.19101	0.03649	0.19315	0.03732
				0.5	0.38918 0.15147	0.3941	0.15532	0.38424	0.14765	0.38945	0.15168	0.39162	0.15337
			0.5	0.3	0.19062 0.03634	0.19536	0.03817	0.18587	0.03455	0.19088	0.03644	0.19297	0.03724
		5	0.3	0.3	0.19071 0.03638	0.19560	0.03827	0.18582	0.03454	0.19098	0.03649	0.19313	0.03731
				0.5	038924 0.15151	0.39415	0.15536	0.38431	0.14770	0.38951	0.15172	0.39167	0.15341
													40

Print ISSN: ISSN 2055-0154(Print), Online ISSN: ISSN 2055-0162(Online)

European Journal of Statistics and Probability

Vol.5, No.3, pp.41-51, June 2017

_Published by European Centre for Research Training and Development UK (www.eajournals.org)

	17	2	0.5 0.3	0.3 0.3	$\begin{array}{c} 0.19051 \\ 0.19345 \end{array}$	0.03630 0.03743	0.19526 0.19830	$\begin{array}{c} 0.03813 \\ 0.03933 \end{array}$	0.18574 0.18858	0.03451 0.03557	0.19077 0.19371	0.03640 0.03753	0.19286 0.19585	$\begin{array}{c} 0.03720 \\ 0.03836 \end{array}$
				0.5	0.39228	0.15389	0.39715	0.15773	0.38739	0.15008	0.39254	0.15409	0.39469	0.15578
			0.5	0.3	0.19255	0.03708	0.19727	0.03892	0.18781	0.03528	0.19281	0.03718	0.19489	0.03799
		5	0.3	0.3	0.19352	0.03745	0.19837	0.03935	0.18865	0.03559	0.19377	0.03755	0.19592	0.03839
				0.5	0.39238	0.15396	0.39724	0.15780	0.38749	0.15015	0.39264	0.15417	0.39478	0.15586
			0.5	0.3	0.19262	0.03710	0.19734	0.03895	0.18787	0.03530	0.19287	0.03720	0.19496	0.03801
10	7	1	0.3	0.3	0.19436	0.03778	0.19919	0.03968	0.18952	0.03592	0.19462	0.03788	0.19675	0.03871
				0.5	0.39353	0.15487	0.39834	0.15868	0.38870	0.15109	0.39378	0.15507	0.39591	0.15675
			0.5	0.3	0.19290	0.03721	0.19762	0.03906	0.18815	0.03540	0.19315	0.03731	0.19524	0.03812

Table 5: The values of BIAS and MSE of the Bayesian estimates of H(t) when t = 3 based on SEL, ABL, LINEX and GE loss functions

n	m	1		κβ	Ĥ	a _{BS}	Ĥ _{BA}	q = -0.5	Ĥ _{BA}	q=0.5	F	BL	Â _{BE}	
п	m		i t	t p	BIAS	MSE	BIAS	MSE	BIAS	MSE	BIAS	MSE	BIAS	MSE
30	23	2	2 0.3	3 0.3	0.16250	0.04092	0.18046	0.04611	0.1 4 446	5 0.036 <mark>3</mark> 9	0.1678	8 0.04208	0.17143	0.0434
				0.5	0.65076	0.43850	0.67290	0.46679	0.62847	0.41106	0.6586	3 0. <mark>44</mark> 810	0.66177	0.4524
			0.5	0.3	0.35963	0.16203	0.38776	0.18080	0.33128	0.14481	0.37229	9 0.16926	0.37363	0.1711
		5	0.3	0.3	0.16297	0.04038	0.18091	0.04563	0.14495	0.03579	0.16833	0.04159	0.17189	0.0429
				0.5	0.65067	0.43775	0.67 <mark>2</mark> 81	0.46608	0.62838	0.41026	0.65854	0.44736	0.66168	0.4517
			0.5	0.3	0.37056	0.17199	0.39834	0.19095	0.34257	0.15455	0.38296	0.17917	0.38439	0.18122
	17	2	0.3	0.3	0.21084	0.05778	0.22769	0.06427	0.19389	0.05186	0.21550	0.05925	0.21922	0.06093
				0.5	0.73882	0.56094	0.75846	0.58932	0.71901	0. <mark>5</mark> 3315	0. <mark>74</mark> 450	0.56945	0.74860	0.57497
			0.5	0.3	0.41184	0.20502	0.43885	0.22552	0.38458	0.18591	0. <mark>4</mark> 2341	0.21253	0.42529	0.21503
		5	0.3	0.3	0.21830	0.05874	0.23488	0.06551	0.20161	0.05252	0.22278	0.06031	0.22655	0.06204
				0.5	0.74176	0.56431	0.76129	0.59271	0.72207	0.53650	0.74786	0.57279	0.75148	0.57835
			0.5	0.3	0.42540	0.21215	0.45192	0.2 <mark>33</mark> 25	0.39864	0.19237	0.43653	0.21993	0.43861	0.22247
10	7	1	0.3	0.3	0.18518	0.06978	0.22602	0.08121	0.14400	0.06213	0.19756	0.07088	0.20529	0.07 <mark>49</mark> 3
				0.5	0.69890	0.53419	0.74707	0.59679	0.65026	0.47639	0.71578	0.55187	0.72263	0.56 <mark>43</mark> 7
			0.5	0.3	0.38481	0.24465	0.44885	0.28282	0.32000	0.2159	0.41452	0.25238	0.41639 0	.26225

CONCLUSION

The Bayes estimators under both symmetric and asymmetric loss functions are obtained. When we compare the loss functions, we find that the ABL loss function is better than SEL, LINEX and GE loss functions in tables 1,3,4, and 5, while find that the ABL and LINEX loss functions are better than SEL and GE loss functions in table 2. _Published by European Centre for Research Training and Development UK (www.eajournals.org)

REFERENCES

- [1] Al-Aboud, F.M. (2009). Bayesian estimations for the extreme value distribution using progressive censored data and asymmetric loss. *International Math. Forum* 4, 1603-1622.
- [2] Al-Bayyati, H.N. (2002). Comparing methods of estimating Weibull failure models using simulation. *Ph.D. Thesis, College of Administration and Economics, Baghdad University, Iraq.*
- [3] Balakrishnan, N. and Sandhu R.A. (1996). Best linear unbiased and maximum likelihood estimation for exponential distributions based on general progressive type-II censored samples. *Sankhya*. 1, 1-9.
- [4] Balakrishnan, N., Cramer, E., Kamps, U. and Schenk, N. (2001). Progressive type II censored order statistics from exponential distributions. *Statist.* 35, 537–556.
- [5] Bernardo, J.M. and Smith, A.F.M. (1994). Bayesian Theory. *Wiley, New York*.
- [6] Calabria, R. and Pulcini, G. (1996). Point estimation under asymmetric loss functions for left truncated exponential samples. *Communications in Statist. Theory and Methods* 25, 585-600.
- [7] Davis, H.T. and Feldstein, M.L. (1979). The generalized Pareto law as a model for progressively censored survival data. *Biometrika* 66, 299–306.
- [8] Figueiredo M.A.T. (2004). Lecture notes on Bayesian estimation and classification. *Instituto de Telecomunicacoes and Instituto Superior Tecnico, Portugal.*
- [9] Gompertz, B. (1825). On the nature of the function expressive of the law of human mortality and on a new model of determining the value of life contingencies. *Philos. Trans. Roy. Soc. London.* 115, 513 585.
- [10] Guilbaud, O. (2001). Exact non-parametric confidence intervals for quantiles with progressive type-II censoring. *Scand. J. Statist.* 28, 699–713.
- [11] Kim C. (2006). On estimating Burr type XII parameter based on general Type-II progressive censoring. *The Korean Communications In Statist. 1, 89-99.*
- [12] Mohie El-Din, M.M., Amein, M.M., El-Attar, H.E. and Hafez, E.H. (2013). Estimation for parameters of Feller-Pareto distribution from progressive type-II censoring and some characterizations. *J. of Probability and Statist. Sciences* 11, 97-108.
- [13] Mohie El-Din, M.M., Sadek, A., Marwa M.Mohie El-Din and Sharawy, A.M. (2017). Characterization for Gompertz distribution based on general progressively type-II right censored order statistics. *International J. of Advanced Statist. and Probability.* 5, 52-56.
- [14] Mohie El-Din, M.M., Sadek, A., Marwa M.Mohie El-Din and Sharawy, A.M. (2017). Characterization for the linear failure rate distribution by general progressively type-II right censored order statistics. *American J. of Theoretical and Applied Statist.* 6, 129-140.