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CARBON EMISSION AND ECONOMIC GROWTH OF SAARC COUNTRIES-A VECTOR AUTOREGRESSIVE (VAR) ANALYSIS

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ABSTRACT: This paper examines the causal relationship between carbon () emissions and economic growth in seven SAARC countries using time series data for the period from 1972-2012. We applied Vector Error Correction Modeling (VECM) approach. We have also applied Augmented Dickey-Fuller (ADF) and Phillips-Perron (P.P) test and Johansen's cointegration approach to check time series properties and cointegration relationship of the variables. Results exhibit a cointegration relationship between environmental pollution and economic growth. Results also show that the estimated coefficients of emissions have positive and significant impacts on GDP in the long run. These results will help the environmental authorities to understand the effects of economic growth on environment for degradation and manage the environmental problems using macroeconomic methods.

KEYWORDS: SAARC, Emission, GDP, Causality, VECM.

INTRODUCTION

South Asian Association for Regional Cooperation (SAARC) consists of eight countries¹ which are characterized by relatively high densities of population, low per capita income and literacy rate, and unplanned use of technology in various sectors that causes environmental degradation.

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Conventional wisdom is that higher economic growth requires huge energy consumption which causes emission of higher level of CO_2 and this in turn deteriorates environmental pollution and threatens the sustainability of environment. Now a day's climate change and global warming have attracted considerable attention worldwide.

Many scholars carried out theoretical and empirical researches on relationship between carbon dioxide emissions and economic growth from the view of EKC hypothesis and decoupling theory. This article will focus on relationship between SAARC carbon dioxide emissions and economic growth during the period of 1972-2012, meanwhile applying Vector Auto Regression (VAR) theory to analyze changes of SAARC environmental pressures in the process of economic growth.

 CO_2 Emissions account for the largest share of total greenhouse gas emissions which are most largely generated by human activities (World Bank, 2007). Rapid increase of CO_2 emissions is mainly the results of human activities due to the development and industrialization over the last decades. It is highly dependent to the energy consumption which is inevitable for economic growth.

Chebbi and Boujelbene (2008), Hatzigeorgiou et al. (2013), Shaari et al. (2012), Ozturk and uddin (2011), Boopen and Harris (2012), Ong and sek (2013), Tiari (2011), Böhm (2011), Wahid et al. (2013), Dantama and Inuwa (2012), Amin (2012), Nain (2013), Dhungel (2008), Muhammad and smile (2012), Jinke and Zhongxue (2011), Noor and Siddiqi (2010), found causal relationship

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¹These countries are Afghanistan, Bangladesh, Bhutan, India, Maldives, Nepal, Pakistan and Sri Lanka.

Published by European Centre for Research Training and Development UK (www.ea-journals.org) between energy consumption, CO_2 emission and economic growth by applying cointegration and vecto error correction econometric model.

McKinesy Global Institute, (2008) analyzed that the successful actions on solving climate change problems should meet at least two conditions, (i) curb the increase of global carbon emissions effectively and (ii) this actions of solving global warming problem should not at the expense of declining economic development and people's living standard. Kaplan et al.(2011) found that the coefficients of the ECT terms for all models are statistically significant implying the long-run bidirectional causal relationship between energy and GDP shows that the higher the level of economic activity the higher the energy consumption and vice versa. The intergovernmental panel on climate change (IPCC, 2007) reported a 1.1 to 6.4 ° c increase of the global temperatures and a rise in sea level of about 16.5 to 53.8 cm by 2100. This would have tremendous negative impact on half of the world's population lives in coastal zones (Lau et al., 2009). In this respect most of the SAARC countries situated in coastal areas and for the global warming it has the vast and negative impact of climate change on SAARC countries.

One of the crucial elements for continuous economic growth, it needed to consumption of more energy that generates huge amounts of CO_2 . Several studies emerged in this regard. Bloch, et al. (2012) found that there is a unidirectional causality running from coal consumption to GDP both in short and long run under supply side analysis and bi-directional causality under demand side analysis between the variables in China. Jalil and Mahmud (2009) found a unidirectional causality running from economic growth to CO_2 emissions in China. Andreoni, and Galmarini (2012)

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Published by European Centre for Research Training and Development UK (www.ea-journals.org) researched the decoupling relationship between economic growth and carbon dioxide (CO_2) emissions in Italian by the way of making a decomposition analysis of Italian energy consumption. Holtz-Eakim and Selden (1995) found that there is a diminishing marginal propensity to emit CO_2 as economies develop. Bhattachryya and ghoshal (2009) analyzed that the inter relationship between the growth rates of CO_2 emissions and economic development is mostly significant for countries that have a high level of CO_2 emissions and pollution. Asafu-Adjaye (2010) found in a study on economic growth and energy consumption in four Asian developing economies that a combination of unidirectional and bidirectional causality between the variables. Hye and Mashkoor (2010) found bidirectional causality between economic growth and environmental sustainability. Apergis and Payne (2009) examined the relationship between CO_2 emissions, energy consumption and output in Central America and they found unidirectional causality from energy consumption and real output to emissions in the short run but there appears bi-directional causality between the variable in the long run.

This study designed to evaluate the causal relationship between CO_2 Emission and GDP growth in SAARC countries applying vector error correction modeling approach covering a period of data from 1972-2012 and suggest some policies to policy makers.

DATA AND THEORETICAL ISSUES

Data

This paper uses annual time series data of real per capita GDP and CO_2 emissions covering the period from 1972 to 2012 for the seven SAARC countries- Bangladesh, Bhutan, India, Maldives,

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Nepal, Pakistan and Sri Lanka. Real per capita GDP is taken as US dollar (\$) and CO_2 emissions variable is metric tons per capita. The data have been obtained from online version of World Development Indicators, the World Bank.

Theoretical Issues

This paper analyses the relationship between the long run causal relationships of economic growth and CO_2 emission in SAARC countries. The hypothesis tests in this paper is whether CO_2 Emission is related to the economic growth. We can express the relationship applying the following functional form between CO_2 emission and economic growth (GDP) as follows:

$$CO_2 = f(GDP) \tag{1}$$

 CO_2 emission and economic growth are likely to have four types long run relationships as i) economic growth can cause CO_2 emission, ii) CO_2 emission can cause economic growth, iii) CO_2 emission and economic growth can simultaneously cause each other and iv) finally CO_2 emission may neither causes economic growth nor does economic growth cause CO_2 emission.

METHODOLOGY

Assessment of Granger causality between the variables and the direction of their causality in a vector error correction framework requires three steps. The first step is to test the nonstationarity property and determine order of integration of the variables, the second step is to detect the existence of long run relationship and the third step is check the direction of causality between the variables.

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Testing for Nonstationarity Property and Order of Integration

Examining the time series properties or nonstationarity properties of the variables is imperative as regression with nonstationary variables provides spurious results. Therefore, before moving further variables must be made stationary. This study applies two unit root tests-the Augmented Dickey Fuller test (Dickey & Fuller, 1979) and Phillips-Perron (Phillips-Perron, 1988) to test whether the variables are nonstationary and if nonstationary the order of integration is the same or not.

Augmented Dickey Fuller (ADF) Test:

The Augmented Dickey-Fuller (ADF) test is used to test for the existence of unit roots and determine the order of integration of the variables. The ADF test requires the equations as follows

$$\Delta y_t = \alpha_0 + \alpha_1 t + \theta y_{t-1} + \sum_{i=1}^m w_i \Delta y_{t-i} + \varepsilon_t$$
(2)

Where, Δ is the difference operator, y is the series being tested, m is the number of lagged differences and ε is the error term.

Phillips-Perron (P.P) Test:

Phillips-Perron (1988) test deals with serial correlation and heteroscedasticity. Phillips and Perron use non parametric statistical methods to take care of serial correlation in the terms with adding lagged difference terms. Phillips-Perron test detects the presence of a unit root in a series. Suppose, y_t is estimating as

$$\Delta y_t = \alpha + \beta t + \rho * y_{t-1} + u_t \tag{3}$$

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Where, the P.P test is the t value associated with the estimated co-efficient of ρ^* . The series is stationary if ρ^* is negative and significant. The test is performed for all the variables where both the original series and the difference of the series are tested for stationary.

Cointegration

We apply Johansen and Juselius (1990) and Johansen (1988) maximum likelihood method to test for cointegration between the series of carbon emission and economic growth. This method provides a framework for testing of cointegration in the context of Vector Autoregressive (VAR) error correction models. The method is reliable for small sample properties and suitable for several cointegration relationships. The cointegration technique uses two tests-the maximum Eigen value statistics and trace statistics in estimating the number of cointegration vectors. The trace statistic evaluates the null hypothesis that there are at most r cointegrating vectors whereas the maximal Eigen value test evaluates the null hypothesis that there are exactly r cointegrating vectors. Let us assume that y_t follows I(1) process, it is an nX1 vector of variables with a sample of t. Deriving the number of cointegrating vector involves estimation of the vector error correction representation:

$$\Delta y_t = \mu_0 + \Pi y_{t-m} + \sum_{i=1}^m \mu_i \Delta y_{t-i} + \varepsilon_t$$
(4)

The long run equilibrium is determined by the rank of Π . The matrix Π contains the information on long run relationship between variables, that is if the rank of Π =0, the variables are not cointegrated. On the other hand if rank (usually denoted by r) is equal to one, there exists one cointegrating vector and finally if 1<r<n, there are multiple cointegrating vectors and there are nXr

Published by European Centre for Research Training and Development UK (www.ea-journals.org) matrices of α and β such that $\Pi = \alpha \beta'$, where the strength of cointegration relationship is measured by α , β is the cointegrating vector and $\beta' y_t$.

The tests given by Johansen and Juselius (1990) are expressed as follows. The maximum Eigenvalue statistic is expressed as:

$$\lambda_{\max} = -T \ln(1 - \hat{\lambda}_{(r+1)})$$
(5)

While the trace statistic is written as follows:

$$\lambda_{trace}(r) = -T \sum_{i=r+1}^{k} \ln(1 - \hat{\lambda}_i)$$
(6)

Where, r is the number of cointegrating vectors under the null hypothesis and $\hat{\lambda}_i$ is the estimated value for the *i*th ordered eigenvalue from the matrix Π . To determine the rank of matrix Π , the test values obtained from the two test statistics are compared with the critical value from Mackinnon-Haug-Michelis (1999). For both tests, if the test statistic value is greater than the critical value, the null hypothesis of *r* cointegrating vectors is rejected in favor of the corresponding alternative hypothesis.

Error Correction Mechanism

The direction of the causality of long run cointegrating vectors in a vector error correction framework can be conducted once the long run causal relationship between the variables is established. Assuming that the variables are integrated of the same order and cointegrated, the following Granger causality test with an error correction term can be formulated:

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$$\Delta Ep_{t} = \eta_{0} + \sum_{i=1}^{n} \alpha i \Delta Ep_{t-i} + \sum_{j=1}^{m} \beta_{j} \Delta GDP_{t-j} + \delta ECT_{t-1} + \varepsilon_{t}$$
(7)

$$\Delta GDP_{t} = \mu_{0} + \sum_{i=1}^{n} \gamma_{i} \Delta GDP_{t-i} + \sum_{j=1}^{m} \lambda_{j} \Delta Ep_{t-j} + \delta ECT_{t-1} + \varepsilon_{t}$$
(8)

Where, ECT is error correction term. This provides the long run and short run dynamics of cointegrated variables towards the long run equilibrium. The coefficient of error correction term shows the long term effect and the estimated coefficient of lagged variables shows the short term effect between the variables.

EMPIRICAL RESULTS

Results of Unit Root Test

The results of the Augmented Dickey Fuller (1981), ADF Stationarity test in levels show that some variables are stationary and some are non-stationary in level form. In the next step of difference form it is found that all the variables are stationary. The results of the stationarity test in levels and in difference form in shown is Table 1.

Table 1:	Augmen	ted Dicke	y-Fuller Un	it Root Tes	t Results						
Level Form					Difference Form						
	With Constant and Trend				With Constant and Trend						
	Statistics	cs Critical Values			Statistics	Critical Values					
Variab		1%	5%	10%		1%	5%	10%			
Bangladesh											
CO_2	-	-	-3.529758	-3.196411	-	-	-	-			
GDP	4.2579	-	-	-	-	-	-3.544284	-3.204699			
	Bhutan										
CO_2	-	-	-3.526609	-3.194611	-	-	-	-			
GDP	0.8132	-	-3.533083	-3.198312	-	-	-	-			
				India							
CO_2	1.0237	-	-3.529758	-3.196411	-	-	-	-			
GDP	2.9502	-	-3.526609	-3.194611	-	-	-	-			
Maldives											
CO_2	-	-	-3.540328	-3.202445	-	-	-	-			
GDP	-	-	-3.536601	-3.200320	-	-	-	-			

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Nepal											
CO_2	-	-	-3.540328 -	-3.202445	-	-	-	-			
GDP	-	-	-3.533083 -	-3.198312	-	-	_	-			
Pakistan											
CO_2	-	-	-3.526609	-3.194611	-	-	-	-			
GDP	-	_	-3.529758	-3.196411	-	-	-	-			
	Sri Lanka										
CO_2	-	-	-3.526609	-3.194611	-	-	-	-			
GDP	4.9199	-	-	-	-	-	-	-			
	49	4.20500	3.526609**	3.194611*	3.712592	4.211868	3.529758*	3.196411**			
		4*		**			*	*			

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The test is conducted using Eviews 7.1

Note: On the base of critical value in Table 1, the * denotes that the rejection of null hypothesis of unit root at 1%, **denotes that the rejection of null hypothesis of unit root at 5% and *** denotes that the rejection of null hypothesis of unit root at 10% level of significance. Here we consider the variables with constant and trend both in level and first difference form.

It is evident from Table that the calculated ADF statistics in respect of Bangladesh CO_2 and GDP are greater than their critical values (denoted by asterisks) in difference form respectively. So in this case the null hypothesis can be rejected. In respect of Bhutan CO_2 and GDP, we found that the calculated ADF statistic is greater than their critical value both in difference and level form respectively. So, null hypothesis can be rejected. For the Indian side we see that the Indian CO_2 and GDP calculated ADF are greater than their critical value both in difference and level form. So, null hypothesis rejected here and so on for Maldives, Nepal, Pakistan and Sri Lanka, it shows that the calculated ADF statistics are greater than their critical value. So, the null hypothesis is rejected and the variables are stationary.

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Table	2: Results	s of Philli	ps-Perron (P.P) Test					
Level form						Difference Form			
Varia	Statistic	Critical	Values		Statistic	Critical Values			
bles	With Consta nt and	1%	5%	10%	With Constant and	1%	5%	10%	
				Banglades		•			
CO_2	-	-	-	-3.196411	-	-	-	-	
GDP	6.0263	-	-	-	-	-	-	-	
				Bhutan		1		I	
CO_2	-	-	-	-3.194611	-	-	-	-	
GDP	0.8132	-	-	-3.198312	-	-	-	-	
	1	I		India		1		I	
CO_2	1.0237	-	-	-3.196411	-	-	-	-	
IGDP	4.4254	-	-	-	-	-	-	-	
			-	Maldives				-	
CO_2	-	-	-	-3.202445	-	-	-	-	
GDP	-	-	-	-3.200320	-	-	-	-	
				Nenal					
CO_2	-	-	-	-3.202445	-	-	-	-	
GDP	-	-	-	-3.198312	-	-	-	-	
	•			Pakistan					
CO_2	-	-	-	-3.194611	-	-	-	-	
GDP	-	-	-	-3.196411	-	-	-	_	
				Sri Lanka					
CO_2	-	-	-	-3.194611	-	-	-	-	
GDP	6.6867	-	-	-	-	-	-	-	

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The test is conducted using Eviews 7.1

Note: On the base of critical value in Table 2, the * denotes that the rejection of null hypothesis of unit root at 1%, **denotes that the rejection of null hypothesis of unit root at 5% and *** denotes that the rejection of null hypothesis of unit root at 10% level of significance. Here we consider the variables with constant and trend both in level and first difference form.

Phillips-Perron Test used to non parametric statistical methods to take care of the serial correlation in the terms without adding lagged difference terms. Table 2 shows the Phillips-Perron (1988) tests

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Published by European Centre for Research Training and Development UK (www.ea-journals.org) results. It is evident from Table 2 that the calculated Phillip-Perron (P.P.) statistics in respect of Bangladesh CO_2 and GDP are greater than their critical values (denoted by asterisks) both in difference and level form. In respect of Bhutan, India, Maldives, Nepal, Pakistan and Sri Lanka, we see that the calculated P.P statistics in respect of CO_2 and GDP are greater than their critical values. So, the null hypothesis can be rejected and the data series are stationary.

Cointegration Results

Cointegration test clarifies that the existence of long run equilibrium relationship among the variables. The cointegration technique is meant to calculate two statistics: Trace (λ_{trace}) statistics and the Maximum Eigen value (λ_{max}) statistics. The estimated results, particularly Maximum Eigen value and Trace statistics are presented in the Table 3 which indicates that the statistics value is greater than the critical value. This means that the hypothesis of no cointegration is rejected and hence they are cointegrated. The Trace statistics and Maximum Eigen value tests indicate that there is one cointegration eqn(s) at 5% level. This means that the variables among environmental pollution (i.e. CO_2 emission) and economic growth (i.e. GDP) have the long run relationships. So, it is clear that there is one linear cointegration eqn(s) for each of the variables that there is one long run relationship and liner deterministic trend among the variables.

More specifically, Table 3 shows that at 5 percent level of significance the likelihood ratios (trace statistics) for the null hypothesis having one (r=1) cointegration (Bangladesh 52.09660, Bhutan 20.14684, India 31.24033, Maldives 30.52002, Nepal 26.51150, Pakistan 35.34613 and Sri Lanka 27.80299) are higher than their respective critical values (Bangladesh 15.49471, Bhutan 15.49471, India 25.87211, Maldives 25.87211, Nepal 25.87211, Pakistan 25.87211 and Sri Lanka 15.49471). At 5% level of significance, the maximum eigenvalue statisticsfor the null hypothesis having one

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Published by European Centre for Research Training and Development UK (www.ea-journals.org) cointegration (Bangladesh 50.89387, Bhutan 19.79190, India 26.51020, Maldives 21.64308, Nepal 21.65528, Pakistan 31.54539 and Sri Lanka 25.86416) are higher than their respective critical values (Bangladesh 14.26460, Bhutan 14.26460, India 19.38704, Maldives 19.38704, Nepal 19.38704, Pakistan 19.38704 and Sri Lanka 14.26460). Hence, according to the likelihood ratio and maximum Eigen value statistics tests, carbon emission and economic growth are cointegrated. Thus, the long run equilibrium relationship among these series is cointegrated.

Table 3: Co-integration Results										
Variable	H0	H1	Trace	5% Critical	Max. Eigen	5% critical	Hypothesis			
				Bangladesh			•			
CO_2 GDP	r=0	r=1	52.09660	15.49471	50.89387	14.26460	Ho: Rejected			
2	r=1	r=2	1.202731	3.841466	1.202731	3.841466	H1: Accepted			
Bhutan										
CO_2	r=0	r=1	20.14684	15.49471	19.79190	14.26460	Ho: Rejected			
2	r=1	r=2	0.354942	3.841466	0.354942	3.841466	H1: Accepted			
India										
CO_{2}	r=0	r=1	31.24033	25.87211	26.51020	19.38704	Ho: Rejected			
2	r=1	r=2	4.730134	12.51798	4.730134	12.51798	H1: Accepted			
				Maldives						
CO_2	r=0	r=1	30.52002	25.87211	21.64308	19.38704	Ho: Rejected			
_	r=1	r=2	8.876940	12.51798	8.876940	12.51798	H1: Accepted			
				Nepal						
CO_2	r=0	r=1	26.51150	25.87211	21.65528	19.38704	Ho: Rejected			
	r=1	r=2	4.856219	12.51798	4.856219	12.51798	H1: Accepted			
Pakistan										
CO_2	r=0	r=1	35.34613	25.87211	31.54539	19.38704	Ho: Rejected			
2	r=1	r=2	3.800743	12.51798	3.800743	12.51798	H1: Accepted			
Sri Lanka										
CO_2	r=0	r=1	27.80299	15.49471	25.86416	14.26460	Ho: Rejected			
2	r=1	r=2	1.938833	3.841466	1.938833	3.841466	H1: Accepted			

Note: The Trace and Max. Eigen value test indicates that there is at least one (1) cointegrating eqn(s) at 5% level of significance. Here ** denotes the rejection of the hypothesis at 0.05 level.

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Results of Error Correction Modeling

Engle and Granger (1987) showed that, if two variables (say X and Y) are individually integrated of order one [i.e. I (I)] and cointegrated then there is possibility of a causal relationship in at least one direction. That means cointegration with I (1) variables indicate the presence of Granger causality but it does not indicate the direction of causality. The vector error correction model is used to detect the direction of causality of long-run cointegrating vectors. Moreover, Granger Representation Theorem indicates how to model a cointegrated series in a Vector Auto Regressive (VAR) format. VAR can be constructed either in terms of level data or in terms of their first differences [I (0)] with the addition of an error correction to capture the short run dynamics.

If the two variables are cointegrated, there must exist an error correction mechanism. This implies that error correction model is associated with the cointegration test. The long term effects of the variables can be represented by the estimated cointegration vector. The adjusted coefficient of error correction term shows the long term effect and the estimated coefficient of lagged variables shows the short term effect. Causality test among the variables are based on Error Correction Model with first difference. Table 4 shows the results of error correction model of the variables.

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Table 4: Results	of Error (Correction	Model							
	Coefficie				Coefficie					
	nt	t	F		nt	t	F			
			Bangla	desh	I		-			
$GDP = f(CO_2)$	0.012022	[0.42823]	1.867654	$CO_2 = f(GDP)$		[50.44211			
Bhutan										
$GDP = f(CO_2)$	0.002749	[0.23656]	0.364334	$CO_2 = f(GDP)$	-	[-	8.089451			
India										
$GDP = f(CO_2)$	-0.002613	[-0.43108]	9.506284	$CO_2 = f(GDP)$	-	[-	17.17979			
			Maldi	ves						
$GDP = f(CO_2)$	-	[-3.72978]	7.365691	$CO_2 = f(GDP)$	-79.42380	[-	5.569285			
			Nepa	al						
$GDP = f(CO_2)$	-	[-	1.160219	$CO_2 = f(GDP)$	-	[-	3.250268			
Pakistan										
$GDP = f(CO_2)$	-0.112020	[-	4.644593	$CO_2 = f(GDP)$	131.6173	[2.041946			
Sri Lanka										
$GDP = f(CO_2)$	0.000134	[0.06242]	0.656019	$CO_2 = f(GDP)$	-	[-	16.65960			

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Note: ** denotes the rejection of the hypothesis at 5% level of significance. The ** values are statistically significant and shows the estimated coefficient of lagged variables. Values in the third brackets are t-statistics.

Table 4 shows the significance of Error Correction Term (ECT) for carbon dioxide (CO_2) emission and economic growth (GDP) of SAARC countries. It is evident from the Table that the error correction term (ECT) is significant for the country Bangladesh, India, Nepal, Bhutan and Sri Lanka in term of GDP, i.e. in these country GDP causes CO_2 for the long term perspective. But in Maldives the ECT is significant in respect of CO_2 emission and for Pakistan we did not find the significance of ECT.

Published by European Centre for Research Training and Development UK (www.ea-journals.org) CONCLUSION

This paper examines the long-run causal relationships between CO_2 emissions and economic growth in SAARC countries during the period of 1972-2012. We apply cointegration and VECM to evaluate the relationship. Empirical results suggest that a long run relationship exist between CO_2 emissions and economic growth in SAARC countries. The application of the cointegration based Granger Causality test found that there is a long run (short run also) relationship between economic growth and CO_2 emissions that is energy consumption granger causes CO_2 emissions and economic growth (GDP). Hence, the long run income elasticity of carbon emissions are greater than the short run income elasticity of carbon emissions, which implies that income (GDP) leads to greater carbon dioxide emissions in the SAARC countries. That is why the significant and positive impact of energy consumption is crucial for economic growth, but the rapid pace of CO_2 emissions requires the adoption of environment friendly developed technology or alternative sources of energy for the protection of environment in seven SAARC countries.

REFERENCES

Amin, S. B., Ferdaus, S.S., and Porna, A.K., (2012), Causal Relationship among Energy Use, CO2 Emissions and Economic Growth in Bangladesh: An Empirical Study, World Journal of Social Sciences, 2, 4, 273 – 290.

Apergis, N., and D. Chapman, (1999), *CO*₂ emissions, energy use, and output in Central America, Energy Policy, 37, 3282-3286 Published by European Centre for Research Training and Development UK (www.ea-journals.org) Andreoni, S., and Galmarini, (2012), Decoupling economic growth from carbon dioxide emissions: A decomposition analysis of Italian energy consumption,

http://dx.doi.org/10.1016/j.energy.

Asafu-Adjaye, J., (2000), The relationship between energy consumption, energy prices and economic growth: Time series evidence from Asian developing countries, Energy Economics, 22, 615-625.

Bhattacharyya, R., and T. Ghoshal, (2009), Economic growth and CO_2 emissions, Environmental Development and Sustainability, http:// DOI 10.1007/s10668-009-9187-2

Bloch, H. et al. (2012), Coal consumption, CO_2 emission and economic growth in China: Empirical evidence and policy responses, Energy Economics, 34, 518-528.

Böhm, D.C., (2011), Economic Growth, Energy Demand and the Environment – Empirical
Insights Using Time Series and Decomposition Analysis, <u>www.weissensee-verlag.de</u>
Boopen, S. and Harris, N., (2012), Energy use, Emissions, Economic growth and Trade: Evidence
from Mauritius, ICTI 2012, ISSN: 16941225, www.b.seetanah@uom.ac.mu

Chebbi, H. E. and Boujelbene, Y., (2008), CO2 emissions, energy consumption and economic growth in Tunisia, 12th Congress of the European Association of Agricultural Economists – EAAE 2008.

Dantama, Y. U., and Inuwa, N., 2012, A Granger-Causality Examination of the Relationship between Energy Consumption & Economic Growth in Nigeria, Journal of Economics, Commerce and Research (JECR) ISSN 2278-4977, 2, 2, 14 – 24

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Dickey, D.A., and W.A. Fuller 1981, Distribution of the estimators for the autoregressive time series with a unit root, Econometrica, 49, 1057-1072.

Dhungel, K. R., (2008), A Causal Relationship Between Energy Consumption and Economic Growth in Nepal, Asia-Pacific Development Journal, 15, 1.

Engle, R.F., and C.W.J. Granger, (1987), Cointegration and error correction-representation, estimation and testing, Econometrica, 55, 251-276.

Grossman, G.M, and A.B. Krueger, (1995), Economic growth and the environment, Quarterly Journal of Economics, 110, 2, 353-377

Hatzigeorgiou, E., Polatadis, H., and Haralambopoulos, D., (2013), Modeling the Relationship among Eenrgy demand, CO2 emission and Economic Development: A Survey for the case of Greece, Global NEST Journal, XX, X, XX-XX.

Hey, Qazi, M.A., Mashkoor, Masood Siddiqui, (2010), Growth and Energy Nexus: An Empirical Analysis of Bangladesh Economy, European Journal of Social Science, 15, 2.

Holtz-Eakin, D. and T.M. Selden, (1995), Stocking the fires? CO_2 Emissions and Economic Growth, Journal of Public Economics, 57, 85-101.

Jalil, A., and S.F. Mahmud, (2009), Environment Kuznets curve for CO_2 emissions: a cointegration analysis, Energy Policy, 37, 5167-5172, http://dx.doi.org/101016/j.enpol.

Jinke, L., and Zhongxue, L., (2011), A Causality Analysis of Coal Consumption and Economic Growth for China and India, Natural Resources, 2, 54-60, http://www.scirp.org/journal/nr

Published by European Centre for Research Training and Development UK (www.ea-journals.org) Lau, L.C., Tan, K.T., and A.R., Mohamed, (2009), A Comparative study of the energy policies in Japan and Malaysia in fulfilling their nation's obligations towards the Kyoto protocol, Energy Policy, 37, 4771-4780. http://dx.doi.org/10.1016/j.enpol

McKinsey Global Company, (2008), The carbon productivity challenge: Curbing climate change and sustaining economic growth. *The McKinsey Quarterly*, 7, 2-33.

Muhammad, S., and Smile, D., (2012), Revisiting the Relationship between Coal Consumption and Economic Growth: Cointegration and Causality Analysis in Pakistan, Applied Econometrics and International Development, 12, 1.

Muhittin Kaplan, M.,Ozturk, I., and Kalyoncu, H., (2011), Energy Consumption and Economic Growth in Turkey : Cointegration and Causal Analysis, Romanian Journal of Economic Forecasting – 2/2011

Nain,Z., Ahmad, W., and Kamaiah, B., (2013), Economic Growth, Electricity Consumption and CO2 Emissions in India: A Disaggregated Causal Analysis.

Noor, S., and M. W Siddiqi, (2010), Energy Consumption and Economic Growth in South Asian Countries: A Co-integrated, International Journal of Human and Social Sciences, 5, 14. Ong, S.M., and sek, S.K., (2013), Interactions between Economic Growth and Environmental Quality: Panel and Non-Panel Analyses, Applied Mathematical Sciences, 7, 14, 687 – 700 Ozturk, I. and uddin, G.S., (2011), Causality Among Carbon Emission, Energy Consumption and Growth in India, Economic Research - Ekonomska Istrazivanja, 25, 3, 753 Phillips-Perron, (1988), Testing for Unit root in Time Series Regression, *Biometrica*, 75, 335-346. http://dx.doi.org/10.1093/biomet/75.2.335 Published by European Centre for Research Training and Development UK (www.ea-journals.org) Shaari, M.S., Hussain, N.E., and Ismail, M.S., (2012), Relationship between Energy Consumption and Economic Growth: Empirical Evidence for Malaysia, <u>www.business-systems-</u> review.org, 2, 1.

The World Bank, (2007), Growth and CO_2 emissions: how do different countries fare,

Environment Department, Washington,

http://wwwds.worldbank.org/external/default/WDSContent/IB/2007/12/05/000020953_2007120

5142250 / Renderd / PDF / 417600 EDP01130 Growth 0 and 0 co 201 PUBLIC1.pdf

The World Bank, (2013), GDP and CO_2 emissions of different countries from 1972-2012,

World Development Indicators Statistics

Tiari, A.K., (2011), Energy Consumption, CO₂ Emissions and Economic Growth: Evidence

from India, Journal of International Business and Economy, 12, 1, 85-122.

Wahid, I.N., Aziz, A.A., and Mustapha, N. H. N., (2013), Energy Consumption, Economic Growth and CO2 emissions in Selected ASEAN Countries, PROSIDING PERKEM VIII, JILID, 2, 2013, 758 – 765